

OefenDeeltentamen 1 Inleiding Financiële Wiskunde, 2011-12

1. Consider a 2-period binomial model with $S_0 = 100$, $u = 1.5$, $d = 0.5$, and $r = 0.25$. Suppose the real probability measure P satisfies $P(H) = p = \frac{2}{3} = 1 - P(T)$.

- (a) Consider an option with payoff $V_2 = \left(\frac{S_1 + S_2}{2} - 105\right)^+$. Determine the price V_n at time $n = 0, 1$.
- (b) Suppose $\omega_1\omega_2 = HT$, find the values of the portfolio process $\Delta_0, \Delta_1(H)$ so that so that the corresponding wealth process satisfies $X_0 = V_0$ (your answer in part (a)) and $X_2(HT) = V_2(HT)$.
- (c) Determine explicitly the Radon-Nikodym process Z_0, Z_1, Z_2 , where

$$Z_2(\omega_1\omega_2) = Z(\omega_1\omega_2) = \frac{\tilde{P}(\omega_1\omega_2)}{P(\omega_1\omega_2)}$$

with \tilde{P} the risk neutral probability measure, and $Z_i = E_i(Z)$, $i = 0, 1$.

- (d) Consider the utility function $U(x) = \ln x$. Find a random variable X (which is a function of the two coin tosses) that maximizes $E(U(X))$ subject to the condition that $\tilde{E}\left(\frac{X}{(1+r)^2}\right) = 30$. Find the corresponding optimal portfolio process $\{\Delta_0, \Delta_1\}$.

Solution (a): The risk neutral measure is given by $\tilde{p} = 2/3 = 1 - \tilde{q}$, and the payoff at time 2 is:

$$V_2(HH) = 82.5, V_2(HT) = 7.5, V_2(TH) = V_2(TT) = 0.$$

Now,

$$V_1(H) = \frac{1}{1.25} \left[\frac{3}{4}(82.5) + \frac{1}{4}(7.5) \right] = 51,$$

$$V_1(T) = \frac{1}{1.25} \left[\frac{3}{4}(0) + \frac{1}{4}(0) \right] = 0,$$

and

$$V_0(H) = \frac{1}{1.25} \left[\frac{3}{4}(51) + \frac{1}{4}(0) \right] = 30.6.$$

Solution (b):

$$\Delta_0 = \frac{V_1(H) - V_1(T)}{S_1(H) - S_1(T)} = 0.51,$$

and

$$\Delta_1(H) = \frac{V_2(HH) - V_2(HT)}{S_2(HH) - S_2(HT)} = 0.5.$$

Solution (c): Notice that

$$Z_i(\omega_1, \dots, \omega_i) = E_i(Z_2)(\omega_1, \dots, \omega_i) = \frac{\tilde{P}(\omega_1, \dots, \omega_i)}{P(\omega_1, \dots, \omega_i)}.$$

Thus,

$$\begin{aligned} Z_2(HH) &= \frac{81}{64}, \quad Z_2(HT) = Z_2(TH) = \frac{27}{32}, \quad Z_2(TT) = \frac{9}{16}, \\ Z_1(H) &= \frac{9}{8}, \quad Z_1(T) = \frac{3}{4}, \quad Z_0 = 1. \end{aligned}$$

Solution (d): The fastest way is to use Theorem 3.3.6. We find that $U'(x) = \frac{1}{x}$ and the inverse I of U' is also given by $I(x) = \frac{1}{x}$. Denoting the Radon Nikodym derivative by $Z = Z_2$, the solution X is given by

$$X = I\left(\frac{\lambda Z}{(1.25)^2}\right) = \frac{(1.25)^2}{\lambda Z}.$$

To find λ , we use the constraint

$$\tilde{E}\left(\frac{X}{(1+r)^2}\right) = E\left(\frac{ZX}{(1+r)^2}\right) = \frac{1}{\lambda} = 30.$$

Hence, $\lambda = 1/30$, and $X = \frac{46.875}{Z}$. That is,

$$X(HH) = 37.04, \quad X(HT) = X(TH) = 55.56, \quad X(TT) = 83.33.$$

The corresponding optimal portfolio can be found using Theorem 1.2.2. Writing $X_2 = X$, the wealth process is given by

$$\begin{aligned} X_1(H) &= \tilde{E}_1\left(\frac{X}{1.25}\right)(H) = \frac{1}{1.25}\left[\frac{3}{4}(37.04) + \frac{1}{4}(55.56)\right] = 33.336, \\ X_1(T) &= \tilde{E}_1\left(\frac{X}{1.25}\right)(T) = \frac{1}{1.25}\left[\frac{3}{4}(55.56) + \frac{1}{4}(83.33)\right] = 50.002, \end{aligned}$$

and

$$X_0 = \frac{1}{1.25}\left[\frac{3}{4}(33.336) + \frac{1}{4}(50.002)\right] \approx 30,$$

as required (small discrepancy due to rounding off errors). The optimal portfolio process is given by

$$\begin{aligned} \Delta_0 &= \frac{X_1(H) - X_1(T)}{S_1(H) - S_1(T)} = -0.16667, \\ \Delta_1(H) &= \frac{X_2(HH) - X_2(HT)}{S_2(HH) - S_1(HT)} = -0.1235, \end{aligned}$$

and

$$\Delta_1(T) = \frac{X_2(TH) - X_2(TT)}{S_2(TH) - S_1(TT)} = -0.5554.$$

2. Consider the N -period binomial model, and assume that $P(H) = P(T) = 1/2$ (we use the same notation as the book). Set $X_0 = 0$, and define for $n = 1, 2, \dots, N$

$$X_i(\omega_1 \dots \omega_N) = \begin{cases} 1, & \text{if } \omega_i = H, \\ -1, & \text{if } \omega_i = T, \end{cases}$$

and set $S_n = \sum_{i=0}^n X_i$, $n = 0, 1, \dots, N$.

- (a) Let $Y_n = S_n^2$, $n = 0, 1, \dots, N$. Show that $E_n(Y_{n+1}) = 1 + Y_n$, $n = 0, 1, \dots, N-1$. Conclude that the process Y_0, Y_1, \dots, Y_N is a submartingale with respect to P .
- (b) Let $Z_n = Y_n - n$, $n = 0, 1, \dots, N$. Show that the process Z_0, Z_1, \dots, Z_N is a martingale with respect to P .
- (c) Let $a > 0$, and define $U_n = a^{S_n} \left(\frac{a^2 + 1}{2a} \right)^{-n}$. Show that the process

$$U_0, U_1, \dots, U_N$$

is a martingale w.r.t. P .

Solution (a): First note that $X_n^2 = 1$ for $n = 1, \dots, n$ and $\tilde{E}_n(X_{n+1}) = E(X_{n+1}) = 0$, this follows from the fact that X_{n+1} is independent from the first n tosses. Furthermore,

$$Y_{n+1} = S_{n+1}^2 = (S_n + X_{n+1})^2 = S_n^2 + 2S_n X_{n+1} + 1.$$

Thus,

$$E_n(Y_{n+1}) = S_n^2 + 2S_n E_n(X_{n+1}) + 1 = S_n^2 + 1 = Y_n + 1,$$

where we used the fact that at time n , S_n is known. Since $E_n(Y_{n+1}) = Y_n + 1 > Y_n$, we have that the process Y_0, Y_1, \dots, Y_N is a submartingale with respect to P .

Solution (b): Using again that $X_n^2 = 1$ for $n = 1, \dots, n$, we have

$$Z_{n+1} = (S_n + X_{n+1})^2 - (n+1) = S_n^2 + 2S_n X_{n+1} - n.$$

Since S_n is known at time n and $\tilde{E}_n(X_{n+1}) = E(X_{n+1}) = 0$, we have

$$E_n(Z_{n+1}) = S_n^2 + 2S_n E_n(X_{n+1}) - n = S_n^2 - n = Z_n.$$

Thus, Z_0, Z_1, \dots, Z_N is a martingale with respect to P .

Solution (c): First, we observe that

$$U_{n+1} = a^{S_n} a^{X_{n+1}} \left(\frac{a^2 + 1}{2a} \right)^{-(n+1)},$$

and

$$E_n(X_{n+1}) = E(X_{n+1}) = a/2 + a^{-1}/2 = \frac{a^2 + 1}{2a}.$$

Since S_n is known at time n , we have

$$E_n(U_{n+1}) = a^{S_n} E_n(a^{X_{n+1}}) \left(\frac{a^2 + 1}{2a} \right)^{-(n+1)} = a^{S_n} \left(\frac{a^2 + 1}{2a} \right)^{-n} = U_n.$$

Thus, U_0, U_1, \dots, U_N is a martingale w.r.t. P .

3. Consider the N -period binomial model, and assume that $P(H) = P(T) = 1/2$ (we use the same notation as the book).

- (a) Assume X_0, X_1, \dots, X_N is a Markov process w.r.t. the risk neutral measure \tilde{P} . Consider an option with payoff $V_N = X_N^2$. Show that for each $n = 0, 1, \dots, N-1$, there exists a function g_n such that the price at time n is given by $V_n = g_n(X_n)$.
- (b) Let X_0, X_1, \dots, X_N be an adapted process on (Ω, P) . Consider the random variables U_1, \dots, U_N on (Ω, P) defined by

$$U_i(\omega_1 \dots \omega_N) = \begin{cases} 1/2, & \text{if } \omega_i = H, \\ -1/2, & \text{if } \omega_i = T. \end{cases}$$

Let $Z_0 = 0$, and $Z_n = \sum_{j=0}^{n-1} X_j U_{j+1}$, $n = 1, 2, \dots, N$. Prove that the process Z_0, Z_1, \dots, Z_N is a martingale w.r.t. P .

- (c) Consider the process U_1, \dots, U_N of part (b), and define

$$S_n = \sum_{i=1}^n U_i, \text{ and } M_n = \min_{1 \leq i \leq n} S_i,$$

voor $n = 1, 2, \dots, N$. Show that the process $(M_1, S_1), \dots, (M_N, S_N)$ is Markov w.r.t. P .

Solution (a): Since X_0, \dots, X_N is a Markov process and $V_N = g_n(X_n)$ with $g_n(x) = x^2$, the result follows from Theorem 2.5.8, and which can be easily proved with backward induction as follows. From the hypothesis, the result is true for N . Assume it is true for $n \leq N$, we will show it is true for $n-1$. Now $V_{n-1} = \tilde{E}_n((1+r)^{-1}V_n)$, and by the induction hypothesis, $V_n = g_n(X_n)$ for some function g_n . Since the process, X_0, \dots, X_n is Markov w.r.t. \tilde{P} , there exists a function h_{n-1} such that

$$\tilde{E}_n(V_n) = \tilde{E}_n(g_n(X_n)) = h_{n-1}(X_{n-1}).$$

Set $g_{n-1} = (1+r)^{-1}h_{n-1}$, we then have

$$V_{n-1} = (1+r)^{-1}\tilde{E}_n(V_n) = g_{n-1}(X_{n-1}).$$

Solution (b): We first observe that $Z_{n+1} = Z_n + X_n U_{n+1}$, and $E_n(U_{n+1}) = E(U_{n+1}) = 0$. Thus,

$$E_n(U_{n+1}) = Z_n + X_n E_n(U_{n+1}) = Z_n.$$

So, Z_0, Z_1, \dots, Z_N is a martingale w.r.t. P .

Solution (c): First note that $M_{n+1} = \min(M_n, S_{n+1})$ and $S_{n+1} = S_n + U_{n+1}$. Let $f : \mathbb{R}^2 \rightarrow \mathbb{R}$ be any function, then

$$f(M_{n+1}, S_{n+1}) = f(\min(M_n, S_{n+1}), S_n + U_{n+1}) = F(M_n, S_n, U_{n+1}).$$

Since M_n and S_n depend on the first n tosses while U_{n+1} is independent of the first n tosses, we have by the independence Lemma that

$$E_n(f(M_{n+1}, S_{n+1})) = E_n(F(M_n, S_n, U_{n+1})) = g(M_n, S_n),$$

where

$$g(m, s) = E(F(m, s, U_{n+1})) = \frac{1}{2} \left(f\left(\min\left(m, s + \frac{1}{2}\right), s + \frac{1}{2}\right) + f\left(\min\left(m, s - \frac{1}{2}\right), s - \frac{1}{2}\right) \right).$$

Hence, $(M_1, S_1), \dots, (M_N, S_N)$ is Markov w.r.t. P .